

# ECON 7010 - MIDTERM 1 2016 - ANSWER KEY

1. B    2. C    3. A    4. C    5. C

6. From the F.O.C. for solving for  $b$  we have:

$$X'Xb - X'y = 0$$

$$X'(Xb - y) = 0$$

$$X'e = 0$$

If the model includes an intercept, then one of the columns of  $X$  (for example the 1<sup>st</sup> column) is a column of 'ones'. So,

$$X'e = \begin{bmatrix} 1 & 1 & \dots & 1 \\ x_{11} & & & \\ \vdots & & & \\ e_n \end{bmatrix} \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix} = \begin{bmatrix} \sum e_i \\ \vdots \\ \vdots \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \end{bmatrix}$$

That is, the first element of  $X'e$  is  $\sum e_i$ , which must be equal to 0 by the F.O.C.

$$7. V(b) = V[(X'X)^{-1}X'y]$$

$$= V[(X'X)^{-1}X'(X\beta + \varepsilon)]$$

$$= V[(X'X)^{-1}X'X\beta + (X'X)^{-1}X'\varepsilon]$$

$$= V[\beta + (X'X)^{-1}X'\varepsilon]$$

$$= \cancel{V(\beta)} + V[(X'X)^{-1}X'\varepsilon]$$

By A.5,  $X$  is non-random, and:

$$V(b) = \cancel{V(\beta)} + (X'X)^{-1}X'V(\varepsilon)X(X'X)^{-1}$$

By A.4,  $V(\varepsilon) = \sigma^2 I$ , and:

$$V(b) = \cancel{V(\beta)} (X'X)^{-1}X'\sigma^2 I X(X'X)^{-1}$$

$$= \sigma^2 (X'X)^{-1}X'X(X'X)^{-1}$$

$$= \sigma^2 (X'X)^{-1}$$

Implicitly, we are also using A.1 and A.2

$$8.a) E(\tilde{\beta}_1) = E \left[ \frac{1}{n} \sum \left( \frac{y_i - \bar{y}}{x_i - \bar{x}} \right) \right]$$

$$= E \left[ \frac{1}{n} \sum \left( \frac{\beta_0 + \beta_1 x_i + \varepsilon_i - \beta_0 - \beta_1 \bar{x} - \bar{\varepsilon}}{x_i - \bar{x}} \right) \right]$$

$$= \frac{1}{n} E \left[ \sum \left( \frac{\beta_1 x_i + \varepsilon_i - \beta_1 \bar{x} - \bar{\varepsilon}}{x_i - \bar{x}} \right) \right]$$

$$= \frac{1}{n} \sum \left( \frac{\beta_1 x_i + E(\varepsilon_i) - \beta_1 \bar{x} - E(\bar{\varepsilon})}{x_i - \bar{x}} \right)$$

$$= \frac{1}{n} \sum \left( \frac{\beta_1 x_i - \beta_1 \bar{x}}{x_i - \bar{x}} \right) = \frac{1}{n} \sum \beta_1 = \beta_1$$

b)  $\tilde{\beta}_1$  is linear since it can be written as a weighted sum of the  $y$  data.

$$\tilde{\beta}_1 = \sum c_i (y_i - \bar{y}),$$

$$\text{where } c_i = \frac{1}{n} \frac{1}{x_i - \bar{x}}$$

$$c) \text{ var}(\tilde{\beta}_1) = \text{var} \left[ \frac{1}{n} \sum \left( \frac{y_i - \bar{y}}{x_i - \bar{x}} \right) \right]$$

$$= \frac{1}{n^2 (x_i - \bar{x})^2} \text{var} \sum (y_i - \bar{y})$$

$$= \frac{1}{n^2 (x_i - \bar{x})^2} \sum [\text{var}(y_i) + \text{var}(\bar{y})]$$

$$= \frac{n\sigma^2 + \sigma^2}{n^2 (x_i - \bar{x})^2}$$

While we could compare the variance above to that of  $\text{var}(b_1)$ , we would find that  $\text{var}(\tilde{\beta}_1) > \text{var}(b_1)$ , due to the Gauss-Markov theorem. The inequality is strict since  $\tilde{\beta}_1$  is not the OLS estimator.

9. The population model is:

$$y = X\beta + \varepsilon. \quad (1)$$

With the addition of the dummy variable, the model becomes:

$$y = X_1\beta_1 + X_2\beta_2 + \varepsilon, \quad (2)$$

where  $X_1\beta_1 = X\beta$  from (1), and where  $X_2 = \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$

The OLS estimator for  $\beta_1$  in (2) is:

$$b_1 = (X_1' M_2 X_1)^{-1} X_1' M_2 y,$$

where  $M_2 = I - X_2 (X_2' X_2)^{-1} X_2'$

$$= I - \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix} (1)^{-1} [1 \ 0 \ \dots \ 0]$$

$$= I - \begin{bmatrix} 1 & 0 & \emptyset \\ \emptyset & \ddots & \emptyset \\ \emptyset & \emptyset & 0 \end{bmatrix} = \begin{bmatrix} 0 & & \emptyset \\ & 1 & \emptyset \\ \emptyset & & 1 \end{bmatrix}$$

Note that  $X_1' M_2 X_1 = X_1' X_1 - X_1' \begin{bmatrix} 1 & 0 & 0 \\ 0 & \ddots & 0 \\ 0 & & 0 \end{bmatrix} X_1$ .

$X_1' M_2 y$  is computed similarly. This just drops the first observation.

Alternatively, write the estimator for  $\beta_1$  as:

$$b_1 = (X_1' M_2' M_2 X_1)^{-1} X_1' M_2' M_2 y$$

$$= \left( \begin{bmatrix} 0 & X_{2k} & X_{nk} \\ \vdots & \vdots & \dots \\ 0 & X_{21} & X_{n1} \end{bmatrix} \begin{bmatrix} 0 & \dots & 0 \\ X_{21} & \dots & X_{2n} \\ \vdots & & \vdots \\ X_{n1} & & X_{nk} \end{bmatrix} \right)^{-1} \begin{bmatrix} 0 & X_{2k} & X_{nk} \\ \vdots & \vdots & \dots \\ 0 & X_{21} & X_{n1} \end{bmatrix} \begin{bmatrix} 0 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

Note that  $M_2 X_1 = \begin{bmatrix} 0 & \dots & 0 \\ X_{21} & \dots & X_{2k} \\ \vdots & & \vdots \\ X_{n1} & & X_{nk} \end{bmatrix}$

$b_1 = \sum_{i=2}^n c_i y_i$ , where  $c_i$  is a non-random constant determined by the  $X$  data.