



7 – Joint Hypothesis Tests

Now that we have **multiple “X” variables**, and **multiple β s**, our **hypotheses might also involve more than one β** .

- We shouldn't use t -tests
- We should use the F -test

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The types of hypotheses we are now considering involve multiple coefficients (β s). For example:

$$H_0 : \beta_1 = 0, \beta_2 = 0 \quad \begin{array}{l} q = 2 \\ \text{multiple hyp.} \end{array}$$

$$H_A : \beta_1 \neq 0 \text{ and/or } \beta_2 \neq 0 \quad \text{not } H_0$$

and

$$H_0 : \beta_1 = 1, \beta_2 = 2, \beta_3 = 5 \quad \begin{array}{l} q = 3 \\ \text{multiple hyp.} \end{array}$$

$$H_A : \beta_1 \neq 1 \text{ and/or } \beta_2 \neq 2 \text{ and/or } \beta_3 \neq 5 \quad \text{not } H_0$$

Note that the null hypothesis is wrong if **any** of the individual hypotheses about the β s are wrong. In the latter example, **if $\beta_2 \neq 2$** then the whole thing is wrong. Hence the use of the “and/or” operator in H_A . It is common to omit all the “and/or” and simply write “not H_0 ” for the alternative hypothesis.

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- A joint hypothesis specifies a value (imposes a restriction) for two or more coefficients
- Use q to denote the number of restrictions ($q = 2$ for 1st example, $q = 3$ for second example)

F -tests can be used for **model selection**. Which variables should we leave out of the model?

- If variables are **insignificant**, we might want to drop them from the model
- Dropping a variable means we hypothesize its β is zero
- Dropping multiple variables at once means all of the associated β s are all zero

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Example: CPS data again

```
summary(lm(wage ~ education + gender + age + experience))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-1.9574	6.8350	-0.286	0.775
education	1.3073	1.1201	1.167	0.244
genderfemale	-2.3442	0.3889	-6.028	3.12e-09 ***
age	-0.3675	1.1195	-0.328	0.743
experience	0.4811	1.1205	0.429	0.668

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 4.458 on 529 degrees of freedom
 Multiple R-squared: 0.2533, Adjusted R-squared: 0.2477
 F-statistic: 44.86 on 4 and 529 DF, p-value: < 2.2e-16

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The results of the above regression make me want to drop age and experience.

This corresponds to the hypothesis:

$$H_0: \beta_3 = 0 \text{ and } \beta_4 = 0$$

$$H_A: \text{either } \beta_3 \neq 0 \text{ or } \beta_4 \neq 0 \text{ or both}$$

Why would we want to drop variables?

- ① simpler is better
- ② improves efficiency (reduces variance)

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We can't use t-tests

$$H_0: \beta_3 = 0 \text{ and } \beta_4 = 0$$

Idea (doesn't work): reject H_0 if either $|t_3| > 1.96$ and/or $|t_4| > 1.96$.

Assume H_0 is true

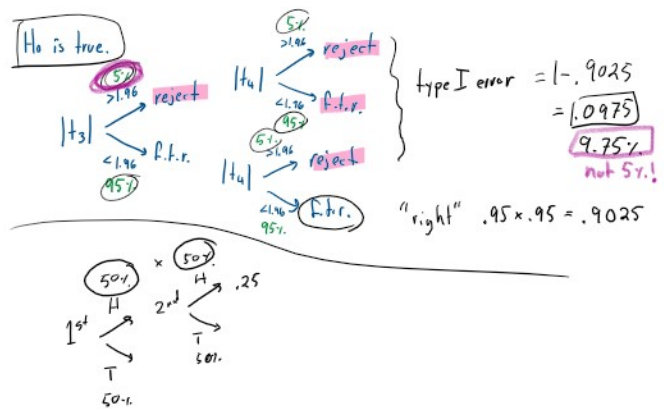
Review: type I error $P_r(\text{reject } H_0 | H_0 \text{ is true}) = \alpha$

Exercise: Assuming that t_3 and t_4 are independent, show that the type I error for the above test is 9.75% (not 5%).

How would you correct this problem? (Bonferroni method - not used in practice)

increase the 1.96
 so that it's harder to reject

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A bigger problem: t_3 and t_4 are likely *not* independent

In the model:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4 + \epsilon$$

- suppose that X_3 and X_4 are *not* independent (e.g. they are correlated)
- then the OLS estimators b_3 and b_4 will be correlated - the formula for b_3 (etc.) involves *all* of the "X" variables (remember OVB)
- then t_3 and t_4 will be correlated!

because $b_3 = f(y, X_1, X_2, X_3, X_4)$ depends on all X
 and because X_3 & X_4 are corr.
 then b_3 & b_4 are correlated

$t_3 = \frac{b_3 - \beta_{3,0}}{s.e.(b_3)}$ $t_4 = \frac{b_4 - \beta_{4,0}}{s.e.(b_4)}$

correlated

Example

Suppose that X_3 and X_4 are *positively* correlated. Consider the null:

$$H_0: \beta_3 = 0 \text{ and } \beta_4 = 0$$

- if b_3 and b_4 are *both positive* (or negative), it's not that big of a deal
- if one is positive and the other negative, that's a big deal

CPS data again

Coefficients:

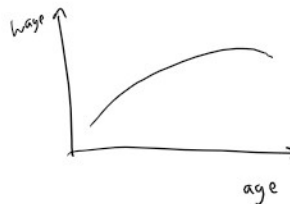
	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-1.9574	6.8350	-0.286	0.775
education	1.3073	1.1201	1.167	0.244
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Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

$$\beta_{age} = \frac{\Delta wage}{\Delta age}; \text{ holding all else constant}$$

- do the *signs of the coefficients* make sense? ✓
- what is the sign of the *correlation* between *age* and *experience*? (+)
- according to the two individual *t*-tests, we fail to reject the null:

$$H_0: \beta_3 = 0 \text{ and } \beta_4 = 0 \quad \left\{ \begin{array}{l} |t_3| < 1.96 \\ |t_4| < 1.96 \end{array} \right\} \Rightarrow \text{f.t.}$$



Let's try the F -test

I'm going to estimate two models:

- One model under the **alternative** hypothesis – we'll call the **unrestricted model** (the β s are allowed to be anything)
- One model under the **null** hypothesis – called the **restricted model**. I get this model by taking the null hypothesis to heart. That is, substitute in the values $\beta_3 = 0$ and $\beta_4 = 0$ into the full model

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Unrestricted model (under H_A): $wage = \beta_0 + \beta_1 educ + \beta_2 gender + \beta_3 age + \beta_4 exper + \epsilon$
`unrestricted <- lm(wage ~ education + gender + age + experience)`

Restricted model (under H_0): $wage = \beta_0 + \beta_1 educ + \beta_2 gender + \epsilon$
`restricted <- lm(wage ~ education + gender)`

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F-test command:

```
anova(unrestricted, restricted)
```

Output (F-stat in blue, p-val in red):

Analysis of Variance Table

Model 1: wage ~ education + gender + age + experience

Model 2: wage ~ education + gender

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	529	10511				
2	531	11425	-2	-914.27	23.007	2.625e-10 ***

Handwritten notes: "OPPOSITE" above the table, "From F-tests" with an arrow pointing to the Pr(>F) column, and "REJECT" with an arrow pointing to the p-value.

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Interpretation? (A big F -stat still means reject)

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A formula for the F-test statistic

- The *F*-test takes into account the correlation between the estimators that are involved in the test
- Note that if the **unrestricted model "fits" significantly better** than the **restricted model**, we should **reject the null**.
- The difference in "fit" between the model under the null and the model under the alternative leads to a formulation of the **F-test** statistic, for testing joint hypotheses.

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The **RSS** is a **measure of fit**:

$$RSS = \sum_{i=1}^n e_i^2$$

where

$$e_i = Y_i - \hat{Y}_i$$

The *F*-test statistic may be written as:

$$F = \frac{(RSS_{restricted} - RSS_{unrestricted})/q}{RSS_{unrestricted}/(n - k_{unrestricted} - 1)} \quad (4)$$

↔ smaller
↔ bigger

*R*² ↑ when we add a variable
 min RSS
 bigger, by
 if small → fail to reject
 if big → reject

where **q = # of restrictions**, **k = # of "X"s**

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Notice that if **the restrictions are true** (if the null is true), **RSS_{restricted} - RSS_{unrestricted}** will be small, and we'll fail to reject.

Another statistic which uses *RSS* is the *R*²:
 $R^2 = 1 - \frac{RSS}{TSS}$ → solve for this → sub. into *F*-test

This gives us another formula for the *F*-test statistic:

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$$F = \frac{(R^2_{unrestricted} - R^2_{restricted})/q}{(1 - R^2_{unrestricted})/(n - k_{unrestricted} - 1)}$$

↑ *H*_A
↑ *H*₀

where:

*R*²_{restricted} = the *R*² for the restricted regression

*R*²_{unrestricted} = the *R*² for the unrestricted regression

$$F = \frac{(R^2_{unrestricted} - R^2_{restricted}) / q}{(1 - R^2_{unrestricted}) / (n - k_{unrestricted} - 1)}$$

where:

$R^2_{restricted}$ = the R^2 for the restricted regression

$R^2_{unrestricted}$ = the R^2 for the unrestricted regression

q = the number of restrictions under the null

$k_{unrestricted}$ = the number of regressors in the unrestricted regression.

The bigger the difference between the restricted and unrestricted R^2 's – the greater the improvement in fit by adding the variables in question – the larger is the F statistic.

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Testing you on the exam

- The F -test statistic can be obtained by comparing the R^2 in the restricted model (H_0 model) and the unrestricted model (H_A model).
- The decision to reject or not depends on whether the F -stat exceeds the (5%) critical value:

q	5% critical value
1	3.84
2	2.00
3	3.00
4	2.00
5	2.00

- These values are only accurate if n is large (we'll always assume this)

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Exercise

Test

$$H_0: \beta_3 = 0 \text{ and } \beta_4 = 0$$

in the model:

$$\text{wage} = \beta_0 + \beta_1 \text{education} + \beta_2 \text{genderfemale} + \beta_3 \text{age} + \beta_4 \text{experience} + \epsilon$$

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Coefficients: *unrestricted*
H_a

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-1.9574	6.8350	-0.286	0.775
education	1.3073	1.1201	1.167	0.244
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 F-statistic: 44.86 on 4 and 529 DF, p-value: < 2.2e-16

Coefficients: *restricted*
H₀

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.21783	1.03632	0.210	0.834
education	0.75128	0.07682	9.779	< 2e-16 ***
genderfemale	-2.12406	0.40283	-5.273	1.96e-07 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 4.639 on 531 degrees of freedom
 Multiple R-squared: 0.1884, Adjusted R-squared: 0.1853
 F-statistic: 61.62 on 2 and 531 DF, p-value: < 2.2e-16

$$F = \frac{(R_u^2 - R_R^2) / q}{(1 - R_u^2) / (n - k_u - 1)} = \frac{(0.2533 - 0.1884) / 2}{(1 - 0.2533) / (534 - 4 - 1)} \approx 23$$